

# 数理統計学から見たThermo-Majorization: 統計モデルの比較と情報スペクトル

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# Thermo Majorization

$\rho_G$ : (Gibbs) state

$$[\rho, \rho_G] = 0, [\sigma, \rho_G] = 0$$

The transition

$$\{\rho, \rho_G\} \rightarrow \{\sigma, \rho_G\}$$

is possible



$(\rho_{ii}/\rho_{G,ii})_i$  majorizes  $(\sigma_{ii}/\rho_{G,ii})_i$

$$\sum_{i=1}^j (\rho_{ii}/\rho_{G,ii}) \geq \sum_{i=1}^j (\sigma_{ii}/\rho_{G,ii}), \forall j$$



$$\text{tr} (\rho - t\rho_G)_+ \geq \text{tr} (\sigma - t\rho_G)_+, \forall t \geq 0$$

$(A)_+$ : positive part of the matrix  $A$

# Thermo-Majorization : History

The term first appears in Horodecki, M + Oppenheim, J (2014)

But had been known since decades ago:

Mathematical Physics : mapping btw. state families  
Uhlmann, Alberiti, Ruch, Schramer, etc ...

- Mathematical Statistics : comparison of “statistical experiments”  
Blackwell, Le Cam, Hajek, Torgersen, etc ...

**Why statisticians are bothered with  
conversion btw state families ?**

# Information Spectrum

A framework of information theory free of stochastic assumptions such as IID, Markov, etc

Classical version : invented by Te-Sun Han (韓太舜),  
with Shannon Award

Quantum version : Hiroshi Nagaoka (UEC)

$$\bar{D}(\{\rho^n\}||\{\sigma^n\}) = \inf\{c; \lim_{n \rightarrow \infty} \text{tr}(\rho^n - e^{nc} \sigma^n)_+ = 0\}$$

$$\underline{D}(\{\rho^n\}||\{\sigma^n\}) = \sup\{c; \lim_{n \rightarrow \infty} \text{tr}(\rho^n - e^{nc} \sigma^n)_+ = 1\}$$

(version by Bowen-Datta)

c.f. Renner, Datta, Bowen give alternative representation of these  
by Smooth Renyi entropy

Computer science, randomness extraction, privacy amplification

## Thermo-Majorization

$$\text{tr} (\rho - t\rho_G)_+ \geq \text{tr} (\sigma - t\rho_G)_+, \forall t \geq 0$$

## Information Spectrum

$$\bar{D}(\{\rho^n\}||\{\sigma^n\}) = \inf\{b; \lim_{n \rightarrow \infty} \text{tr}(\rho^n - e^{nb}\sigma^n)_+ = 0\}$$

$$\underline{D}(\{\rho^n\}||\{\sigma^n\}) = \sup\{b; \lim_{n \rightarrow \infty} \text{tr}(\rho^n - e^{nb}\sigma^n)_+ = 1\}$$

If you are not xxxxxx,

Should notice some relations btw these....

# Plan of the talk:

## Comparison of statistical experiments

- general theory, historical contexts
- 2-states case
- relation to information measures
- quantum states

## Information spectrum (mainly hypothesis test)

- upper & lower divergence rate (classical, quantum)
- hypothesis test  
(- resolvability)
- relation to smooth Renyi entropies

## Asymptotic 2-states conversions

- sufficient condition by information spectrum
- characterization of quantum relative entropy (KL-divergence)

# Comparison of statistical experiments

# References

Blackwell, D. and Girshick, M. A.

*Theory of Games and Statistical Decisions* (1954)

LeCam, L. ( many works since 60's)

*Asymptotic Methods in Statistical Decision Theory*, Springer (1986)

Torgersen, E. *Comparison of Statistical Experiments*,

Cambridge University Press(1991)

Goel P K (**concise and comprehensive review**)

“When is one experiment ‘always better than another?’”

*The Statistician* 52 , Part 4 , pp. 515–537 (2003)

# Statistical decision problems

$x \in X$  : (array of) data       $d(x) \in D$  : decision

$x$  obeys probability distribution  $p_\theta(x)$ ,  $\theta \in \Theta$  unknown

$g_\theta(d) \in \mathbf{R}$ : the gain of the decision  $d$  when the data source is  $p_\theta$

$X$  等は各自分かりやすいので考えて  
(理論は極限まで一般化されてるが気にしないで)

$E = \{p_\theta; \theta \in \Theta\}$  is called an experiment or a statistical model

Optimize  $d$  to maximize the expectation

$$G_\theta(d, E) := E_\theta g_\theta(d(X)) = \sum_{x \in X} g_\theta(d(x)) p_\theta(x)$$

under various setting min-max/average about  $\theta$ /constrained

# Some notes

## Settings are fairly general

$X$  : measurable space  $D$  : topological space, with Baire  $\sigma$ -field,  
 $d$ : measurable

$P_\theta$ : prob measure, may not be majorized by a common measure

$\Theta$ : a set

## $g$ is usually subject to some reasonable conditions

$g_\theta$ : upper-semi continuous/continuous

+ bounded from above/bounded

## In general, randomized decision is also allowed

ここでは記述が面倒だから避けるが、どっか破たんしてたらごめんなさい

## Usually,

“loss “ is minimized, rather than “gain” is maximized

大人の事情があるので、ここではゲイン。

# Example of Statistical Decisions

- Ex 1**  $x$  : past/present data about economics  
 $\theta$  : parameters determining the market price  
 $d$  : your strategy of selling/buying  
 $g_{\theta}(d)$  : how much you earn
- Ex 2**  $\theta \in \mathbf{R}^k$   
 $d$  : your estimate of  $\vartheta$   
 $g_{\theta}(d) = -||d - \theta||^2$  : how close was your guess
- Ex 3**  $\theta \in \{0,1\}$   
 $d$  : your guess about  $\theta \in \{0,1\}$   
 $g_{\theta}(d) = 1$  if you are right, otherwise =0

# On arbitrariness of $g$

**Ex 2**  $\theta \in \mathbf{R}^k$

$d$  : your estimate about the true value of parameter

$g_\theta(d) = -\|d - \theta\|^2$  : how close was your guess

No overwhelming reason to prefer square-error to other error measures

# Statistical Decision Problems : Optimization ?

Life is so hard that **no decision is uniformly optimal usually** :

if  $G_\theta(d, E) := \sum_x g_\theta(x)p_\theta(x)$  very good,  $G_{\theta'}(d, E)$  is poor

**Ex 3**  $\theta \in \{0,1\}$   $d$  : your guess about  $\theta \in \{0,1\}$   
 $g_\theta(d) = 1$  if you are right, otherwise =0

If  $d(x) = 0$  irrespective of  $x$ ,  $G_0(d, E) = 1$  but  $G_1(d, E) = 0$

**1** Maximize the average about  $\theta$  (Bayesian gain):

$$G_\pi(d, E) := \sum_{\theta} \pi_{\theta} G_{\theta}(d, E)$$

**2** Maximize the worst case about  $\theta$  (min-max):

$$G(d, E) = \inf_{\theta \in \Theta} G_{\theta}(d, E)$$

**3** Maximize e.g.  $G_{\theta_1}(d, E)$ , subject to the constrain, e.g.

$$G_{\theta}(d, E) \geq c, \forall \theta \neq \theta_1$$

**And so on ...**

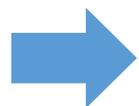
# “optimal decision” ?

- depends on  $G$
- depends on min-max or Bayesian gain, or others

Life is so hard.

You have to solve them one by one.

But, there is a hope ....



## “Comparison of statistical experiments”

- asymptotic theory      存在感こっちのほうが圧倒的で  
    LAN, bounds on cloning etc
- experimental design

でも説明しやすいのはこっちなので・・・

# Experimental design

$$E_F = \{F(f_\theta); \theta \in \Theta\}$$

Given  $\{f_\theta; \theta \in \Theta\}$ , can move  $F$  (experimental design)

Nice if one can say  $\{F(f_\theta); \theta \in \Theta\}$  is more informative than the other

**Ex**

$$x = Ff_\theta + \xi, \quad \xi \sim N(0, \sigma I_k)$$

$f_\theta$ :  $l$ -dim vector,  $x$ :  $k$ -dim vector,  $k \leq l$ ,  $F$ :  $k \times l$  real matrix

Optimize  $F$  obtain most “informative”  $p_\theta(x) = N(Ff_\theta, \sigma I_k)$

If  $N(Ff_\theta, \sigma I_k)$  is better than  $N(F'f_\theta, \sigma I_k)$ , use  $F$  rather than  $F'$

# Comparison of statistical experiments

$$e = (e_\theta)_{\theta \in \Theta} \quad (\text{普通は}\varepsilon\text{を使うが、}\varepsilon\text{ほかのところでも使うので})$$

**DEF**

$E$  is  $e$ -efficient relative to  $E'$ ,  $E \geq_e E'$

$$\Leftrightarrow \forall d' \exists d \forall \theta \in \Theta G_\theta(d, E) \geq G_\theta(d', E') - \frac{1}{2} e_\theta$$

holds for any  $g$  with  $0 \leq g_\theta \leq 1$ , on any  $D$

どんな $E'$ 上のdecision  $d'$ に対しても、それを上回る $E$ 上のdecision  $d$ がある ( $e$ 程度のマージンで)。しかも、それが任意の決定空間と利得で。

$$E \sim_e E' \Leftrightarrow E \geq_e E' \ \& \ E \leq_e E' \quad \begin{array}{l} E' \text{で最適化考えても} \\ E \text{と大体同じ} \end{array}$$

$$\text{If } e_\theta = 0, \forall \theta \in \Theta, \quad E \geq E' \quad E \sim E'$$

“ $\sim$ ” is an equivalence relation

$$\text{Distance measure } \Delta(E, E') := \min \{ \epsilon ; E \sim_e E', e_\theta = \epsilon \}$$

“ $\geq$ ” is partial order

# Randomization criteria (Blackwell, Le Cam)

**Thm**  $E \geq_e E'$

$\Leftrightarrow \exists \Lambda$  : transition probability

$$\forall \theta \in \Theta, \|\Lambda(p_\theta) - p'_\theta\|_1 \leq e_\theta$$

$$\Leftrightarrow \sup_d \sum_{\theta \in \Theta} \pi_\theta G_\theta(d, E) \geq \sup_d \sum_{\theta \in \Theta} (\pi_\theta G_\theta(d, E') - \frac{1}{2} e_\theta)$$

holds for any  $\pi_\theta$  (non-zero at most finitely many points)  
for any  $g$  with  $0 \leq g_\theta \leq 1$ , on any  $D$

This is the reason why statisticians are interested in thermo-majorization-type conditions

# Ex. Comparison of linear normal experiments

**Ex**  $x = Ff_\theta + \xi, \quad \xi \sim N(0, \sigma I_k)$

$f_\theta$ : l-dim vector,  $x$ : k-dim vector,  $k \leq l$ ,  $F$ :  $k \times l$  real matrix

Hansen, O. and Torgersen, E. : *The Annals of Statistics* 1974, Vol. 2, No. 2, 367-373

$$\{N(Ff_\theta, \sigma I_k); \theta \in \Theta\} \geq \{N(F'f_\theta, \sigma I_{k'}); \theta \in \Theta\}$$

$$\Leftrightarrow FF^T - F'F'^T \geq 0$$

( if  $\sigma$  is unknown,  $k - k' - \text{rank}(FF^T - F'F'^T) \geq 0$ , in addition)

many extensions, e.g., about covariance matrix, also computation of  $\Delta$

## Quantum

$\rho_\theta$ : Gaussian state  $E_\theta(\text{tr} \rho_\theta X_1, \text{tr} \rho_\theta P_1, \dots) = Ff_\theta$ ,

[M 2010]

covariance =  $\Sigma$   $J$ : matrix defining CCR

$$\{\rho_\theta; \theta \in \Theta\} \geq^q \{\rho_{\theta'}; \theta' \in \Theta\}$$

$$\Leftrightarrow \exists C, S, \Sigma_\omega \quad C^T J F = J F', \quad \Sigma' = S^T \Sigma_\omega S + C^T \Sigma C, \quad \Sigma_\omega + iJ \geq 0$$

## Ex Statistical Inference on Unital qubit channels

$\Lambda_\theta$ : a unital channel with the unknown parameter  $\theta$   $\Lambda_\theta(I) = I$   
 $F$ : input state + measurement



Many papers, each dealing with each own setting

[Fujiwara02] [Fujiwara03][Sacchi05-1][Sacchi05-2][Sacchi05-3] etc

But all of them say maximally entangled state is optimal

**Def**  $\{\rho_\theta\}_{\theta \in \Theta} \geq^c \{\rho'_\theta\}_{\theta \in \Theta}$   
 $\Leftrightarrow \forall M' \exists M \{P_{\rho_\theta}^M\}_{\theta \in \Theta} \geq \{P_{\rho'_\theta}^M\}_{\theta \in \Theta}$

**Fact** [M2013]

$\Phi$ : max-ent state

$$\forall \rho \{\Lambda_\theta \otimes I(\Phi)\}_{\theta \in \Theta} \geq^c \{\Lambda_\theta \otimes I(\rho)\}_{\theta \in \Theta}$$

# Local asymptotic normality (LAN) $\theta \in R$

- Once upon a time, A. Fisher insisted that maximally likelihood estimate (MLE, 最尤推定量) should be the best estimate, and gave a proof, using “ $\doteq$ ”, “ $\sim$ ”, which would satisfy most of us, but not mathematicians
- Counter example found out, even for Gaussian distributions (Hodge estimator etc)
- Introduction of regularity conditions, on models and estimators
- To simplify the argument, invented was LAN ...

$$\log p_{\theta + \frac{h}{\sqrt{n}}}^n(x^n) - \log p_{\theta}^n(x^n) \sim (J_{\theta})^{-1} \dot{\ell}_{\theta}^n(x^n) h + h^2 J_{\theta}^{-1}$$

If  $\theta \rightarrow \sqrt{p_{\theta}(\cdot)}$  has the **first** derivative ( as a function onto  $L^2$ )

# Local asymptotic normality (LAN) $\theta \in R$

$$E_{\theta}^n := \left\{ p_{\theta + \frac{h}{\sqrt{n}}}^n(x^n); h \in R \right\} \quad E_{\theta} := \left\{ N(h, J_{\theta}^{-1}) \right\}_{h \in R} \quad \theta \in R$$

**Thm**

If  $\theta \rightarrow \sqrt{p_{\theta}(\cdot)}$  has the **first** derivative (as a function on  $L^2$ )

And some mild conditions on  $p_{\theta}$

$$\Delta(E_{\theta}^n, E_{\theta}) \rightarrow 0, \text{ uniformly } \forall \theta \text{ in arbitray compact subset of } \Theta$$

Estimation of  $\theta$  on  $\{p_{\theta}^n\}$  reduces to the one on  $\{N(h, J_{\theta}^{-1})\}_{h \in R}$

If you believe  $x (\sim N(h, J_{\theta}^{-1}))$  is the best estimate of the mean  $h$ , should also believe that MLE is the best

反例は全てガウスでもなので、ガウスに落としてから排除するよう条件つける

Shiryaev, Spokoiny, Statistical Experiments and Decisions: Asymptotic Theory, World Scientific 2000

Ibragimov. Khasminskii, Statistical Estimation: Asymptotic Theory, Springer, 1981

# LAN の面白いところ

- First reduce to the problem to the estimation of the mean of Gaussian random variable, which is easy
- Gaussian shift model  $\{N(h, J_\theta^{-1})\}_{h \in \mathbb{R}^k}$  functions as a “standard form” of parameter family
  - when not iid, other “standard forms”, such as LAMN etc
  - non-parametric

こういった、「問題の帰着」は情報科学だといろんなところに現れる  
計算量理論 (comparison of hardness by reduction)  
量子情報のエンタングルメントの理論 (resource theory)

$$|\Theta| = 2, \quad \mathbf{E} = \{p_0, p_1\}$$

**Thm**

[Torgersen70]

To check  $E \geq_e E'$ , only have to check  
Binary decision problem,  $d \in D = \{0,1\}$

$$\Leftrightarrow \sup_d \sum_{\theta \in \Theta} \pi_\theta G_\theta(d, E) \geq \sup_{d'} \sum_{\theta \in \Theta} (\pi_\theta G_\theta(d', E') - \frac{1}{2} \pi_\theta e_\theta)$$

holds for any  $\pi_\theta$

for any  $g$  with  $0 \leq g_\theta \leq 1$ , on  $D = \{0,1\}$

$$\text{w.l.g., } \pi_0 = 1, \pi_1 = t, \quad g_1(d) = 1 - g_0(d),$$

$$|\Theta| = 2, \quad \mathbf{E} = \{p_0, p_1\}$$

**Thm**

[Torgersen70]

To check  $E \geq_e E'$ , one only have to check  
Binary decision problem,  $d \in D = \{0,1\}$

$$\Leftrightarrow \sup_d \sum_{\theta \in \Theta} \pi_\theta G_\theta(d, E) \geq \sup_{d'} \sum_{\theta \in \Theta} (\pi_\theta G_\theta(d', E') - \frac{1}{2} \pi_\theta e_\theta)$$

holds  $\pi_0 = 1, \pi_1 = t,$

for any  $g$  with  $0 \leq g_\theta \leq 1, \quad g_1(d) = 1 - g_0(d),$

$$\begin{aligned} \sup_d \sum_{\theta \in \Theta} \pi_\theta G_\theta(d, E) &= \sup_d \sum_x g_0(d(x)) [p_0(x) - t p_1(x)] + t \\ &\leq \sum_x (p_0(x) - t p_1(x))_+ + t \quad \text{"=" if } g_0(0) = g_1(1) = 1 \end{aligned}$$

$$|\Theta| = 2, \quad \mathbf{E} = \{p_0, p_1\}$$

**Thm**

[Torgersen70]

To check  $\mathbf{E} \geq_e \mathbf{E}'$ , one only have to check  
Binary decision problem,  $d \in D = \{0,1\}$

$$\mathbf{E} \geq_e \mathbf{E}' \Leftrightarrow$$

$$\forall t \geq 0 \quad \sum_x (p_0(x) - t p_1(x))_+ \geq \sum_x (p_0'(x) - t p_1'(x))_+ + e_0 - t e_1$$

Thermo-majorization with error term

When  $e_\theta = 0$ , thermo-majorization

... had been known since almost half-a-century ago

# Relation to f-divergence

$$\begin{aligned}\sup_{d(\cdot)} \sum_{\theta=0,1} \pi_{\theta} G_{\theta}(d, \mathbb{E}) &= \sup_{d(\cdot)} \sum_{\theta=0,1} \pi_{\theta} \sum_x g_{\theta}(d(x)) p_{\theta}(x) \\ &= \sup_{d(\cdot)} \sum_x \sum_{\theta=0,1} \pi_{\theta} g_{\theta}(d(x)) p_{\theta}(x) = \sum_x \sup_d \sum_{\theta=0,1} \pi_{\theta} g_{\theta}(d) p_{\theta}(x) \\ &= \sum_x p_1(x) f\left(\frac{p_0(x)}{p_1(x)}\right) =: D_f(p_0 || p_1)\end{aligned}$$

$$f(z) := \sup_d \{\pi_0 g_0(d)z + \pi_1 g_1(d)\}$$

f is convex and lower semi-continuous

Any such f can be written in above form (use Legendre transform)

$f(z) = z \log z$  : KL-divergence (or relative entropy)

$z^s$  : Tsallis-Renyi-like quantity

$|1 - z|$  : L1-norm

$(1 - tz)_+$  : Thermo-Majorization

# Relation to f-divergence

## Thm

For any  $g$  and  $\pi$ , there is  $f$  s.t.

$$\sup_{d(\cdot)} \sum_{\theta=0,1} \pi_{\theta} G_{\theta}(d, E) = D_f(p_0 || p_1) := \sum_x p_1(x) f(p_0(x)/p_1(x))$$

And vice versa !

**Any f-divergence is optimized Bayes gain**

$$E \geq_e E' \iff D_f(p_0 || p_1) \geq D_f(p_0' || p_1')$$

holds for any lower semicont. convex function  $f$

## Lem

[CohenKempermanZbaganu98]

Any lower-semicontinuous convex  $f$  can be :

$$f(z) = a + bz + \int (z - t)_+ d\mu(t)$$

Combination of above facts leads to another proof of Thermo-Majorization

## fact

- (i)  $D_f(tp_0 || tp_1) = tD_f(p_0 || p_1)$
- (ii)  $(p_0, p_1) \rightarrow D_f(p_0 || p_1)$  is convex
- (iii)  $D_f(p_0 || p_1) \geq D_f(\Lambda(p_0) || \Lambda(p_1))$   $\Lambda$ : transition probability

# Quantum version I

$$E = \{\rho_\theta\}_{\theta \in \Theta} \quad G_\theta(\Lambda, E) := \text{tr} g_\theta \Gamma(\rho_\theta)$$

$g_\theta$ : bounded self-adjoint, on  $H_D$

$\Gamma$ : Completely positive trace preserving map

**DEF**

$E \geq_e^q E'$   $E$  is  $e$ -defficient relative to  $E'$ ,

$$\Leftrightarrow \forall \Gamma' \exists \Gamma \forall \theta \in \Theta \quad G_\theta(\Gamma, E) \geq G_\theta(\Gamma', E') - \frac{1}{2} e_\theta$$

holds for any  $g$  with  $0 \leq g_\theta \leq I$ , on any  $H_D$

**Thm**

[M10]  $E \geq_e E'$

$$\Leftrightarrow \exists \Lambda : \text{CPTP map}$$

$$\forall \theta \in \Theta, \quad \|\Lambda(p_\theta) - p'_\theta\|_1 \leq e_\theta$$

$$\Leftrightarrow \sup_{\Gamma} \sum_{\theta \in \Theta} \pi_\theta G_\theta(\Gamma, E) \geq \sup_{\Gamma} \sum_{\theta \in \Theta} (\pi_\theta G_\theta(\Gamma, E') - \frac{1}{2} e_\theta)$$

holds for any  $\pi_\theta$  (non-zero at most finitely many points)  
for any  $g$  with  $0 \leq g_\theta \leq I$ , on any  $H_D$

# Quantum version II

**Def**

$$E \geq_e^c E'$$

$$\Leftrightarrow \forall M' \exists M \{P_{\rho_\theta}^M\}_{\theta \in \Theta} \geq_e \{P_{\rho_{\theta'}}^M\}_{\theta \in \Theta}$$

**Fact**

$$E \geq_e^c E' \Leftarrow \exists \Lambda : \text{positive (may not be CP) trace preserving map} \\ \forall \theta \in \Theta, \|\Lambda(p_\theta) - p'_{\theta}\|_1 \leq e_\theta$$

$\Rightarrow$  is not true. Counter example by [M13]

No good necessary and sufficient condition for  $\geq_e^c$

$$E \geq_e^c E' \Leftarrow E \geq_e^q E'$$

The opposite not true. But if  $e=0$ , some good relation.

**Thm**

[Buscemi 12]

$$\forall E_0 \quad E \otimes E_0 \geq_e^c E' \otimes E_0 \Leftrightarrow E \geq_e^q E'$$

# Quantum local asymptotic normality (LAN)

It had been noted that  $\rho_\theta^{\otimes n}$  and its tangent space can be approximated by Gaussian states, showing the achievable lower bound to asymptotic error bound of asymptotically unbiased estimators

[M 98] [HayashiM 2002] [Hayashi 2002]

$$E_\theta^n := \left\{ \rho_{\theta + \frac{h}{\sqrt{n}}}^{\otimes n} ; h \in \mathbf{R}^k \right\} \quad E_\theta: \text{Gaussian shift (multi-mode)}$$
$$h \in \mathbf{R}^k$$

**Thm**

[Kahn Guta 2005]  $\dim H < \infty$

$$\Delta(E_\theta^n, E_\theta) \rightarrow 0, \text{ uniformly } \forall \theta$$

$$|\Theta| = 2, \quad \mathbf{E} = \{\rho_0, \rho_1\}$$

**Thm**

[AlbertiUhlmann85]

Suppose  $\dim H = 2$

$$E \geq^q E' \quad \Leftrightarrow \quad E \geq^c E'$$

$$\Leftrightarrow \quad \forall t \geq 0, \quad \text{tr}(\rho_0 - t \rho_1)_+ \geq \text{tr}(\rho_0' - t \rho_1')_+$$

[M10] [Jencova10]

Suppose  $[\rho_0, \rho_1] = 0$

$$E \geq_e^c E' \quad \Leftrightarrow$$

$$\forall t \geq 0, \quad \text{tr}(\rho_0 - t \rho_1)_+ \geq \text{tr}(\rho_0' - t \rho_1')_+ - \frac{1}{2}(e_0 - t e_1)$$

# Relation to quantum divergence

$D^Q(\rho_0||\rho_1) := \sup_{\Gamma} \sum_{\theta=0,1} \pi_{\theta} \text{tr} g_{\theta} \Gamma(\rho_{\theta})$  satisfies

(i)  $D^Q(t\rho_0||t\rho_1) = D^Q(\rho_0||\rho_1)$

(ii)  $(\rho_0, \rho_1) \rightarrow D^Q(\rho_0||\rho_1)$  is convex

(iii)  $D^Q(\rho_0||\rho_1) \geq D^Q(\Lambda(\rho_0)||\Lambda(\rho_1))$ ,  $\Lambda$ : CPTP

Also, any function satisfying above three is written in the RHS form

If restricted to commutative ops,  $D^Q(p_0||p_1) = D_f(p_0||p_1)$ , for some  $f$

$D^Q$ : quantum version of f-divergence

$$D_f^{\max}(\rho_0||\rho_1) \geq D^Q(\rho_0||\rho_1) \geq D_f^{\min}(\rho_0||\rho_1),$$

**Fact**

$$E \geq^q E' \Leftrightarrow$$

$D^Q(\rho_0||\rho_1) \geq D^Q(\rho_0'||\rho_1')$  holds for all  $D^Q$  with above conditions

Classical and Quantum Information Spectrum,

Especially on divergence rate

# Classical Information Spectrum

Founded by Te-Sun Han (韓太舜)

Novel frame work of information theory :  
no need to assume iid, memoryless, Markov, Ergodic etc.

What is done:

Rewrite the solutions of information theoretic problems into “spectrum formulas”, which is still abstract.

Thus, if more concrete form is necessary, have to evaluate them further.

Yet, information theoretic part finishes at this stage: the rest of the job is mathematical.

Also, abstract “spectrum formulas” help understanding relations btw various information theoretic problems.

As such, if you want to appreciate its full value, have to learn various aspects of information theory.

Here treat only hypothesis test

# Hypothesis test and divergence rate

$E^n = \{p_0^n, p_1^n\}$  : Guess which from the data  $x^n \sim p_\theta^n$

$P^n\{1|0\}$ : Prob of choosing 1 while 0 is true

$P^n\{0|1\}$ : Prob of choosing 0 while 1 is true

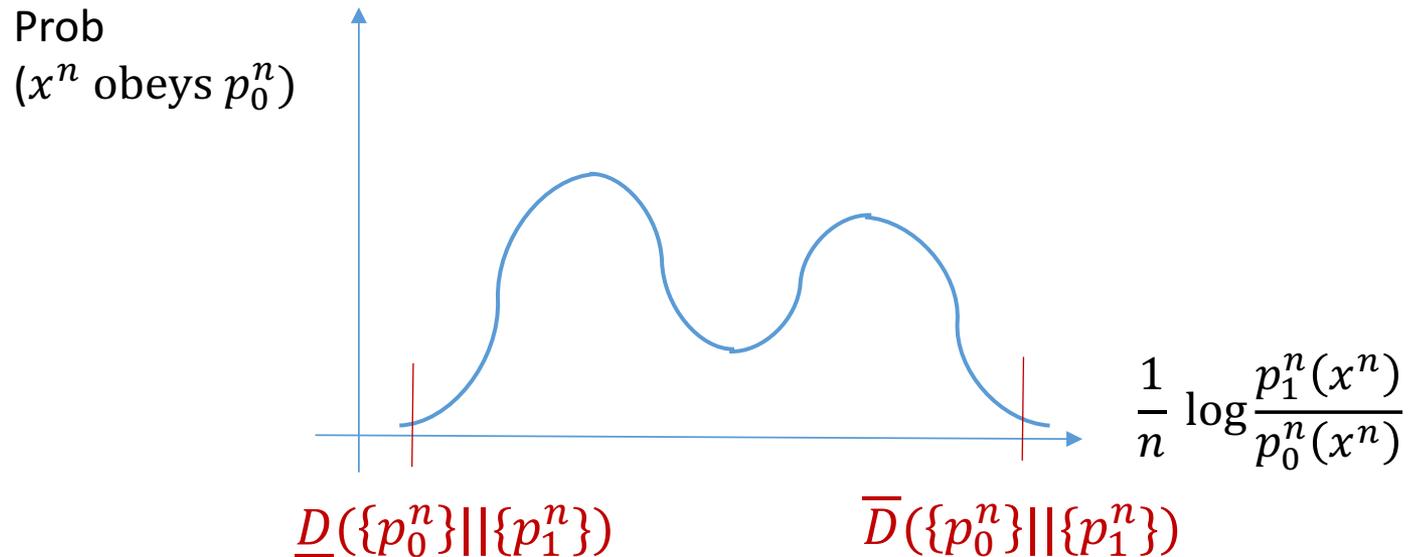
Make  $P^n\{0|1\}$  small, while keeping  $P^n\{1|0\}$  reasonably small

$$e^{-nb} \sim P^n\{0|1\}, \quad b = \limsup_{n \rightarrow \infty} \frac{1}{n} \log P^n\{0,1\}$$

1.  $\sup \{b ; P^n\{1|0\} \rightarrow 0\}$

2.  $\sup \{b ; \limsup P^n\{1|0\} < 1\}$

# Divergence rate and hypothesis test



When iid  $\frac{1}{n} \sum \log \left( \frac{p_1(X_i)}{p_0(X_i)} \right) \rightarrow \sum_x p_0(x) \log \frac{p_0(x)}{p_1(x)} = D(p_0 || p_1)$

$$e^{-nb} \sim P^n\{0 | 1\}, \quad b = \limsup_{n \rightarrow \infty} \frac{1}{n} \log P^n\{0, 1\}$$

$$\sup \{b ; P^n\{1|0\} \rightarrow 0\} = \underline{D}(\{p_0^n\} || \{p_1^n\})$$

$$\sup \{b ; \limsup P^n\{1|0\} < 1\} = \bar{D}(\{p_0^n\} || \{p_1^n\})$$

# Quantum version

By Hirohoshi Nagaoka (長岡浩司) in about 1998

普通、 $\log p/q$  のオペレーター版を考えたいが、そこにハマらなかった  
仮説検定の最適化の途中経過で何がポイントかを考えた。「比の非可換版」

$$\bar{D}(\{\rho_0^n\}||\{\rho_1^n\}) = \inf\{b; \lim_{n \rightarrow \infty} \text{tr}(\rho_0^n - e^{nb} \rho_1^n)_+ = 0\}$$

$$\underline{D}(\{\rho_0^n\}||\{\rho_1^n\}) = \sup\{b; \lim_{n \rightarrow \infty} \text{tr}(\rho_0^n - e^{nb} \rho_1^n)_+ = 1\}$$

1. When commutative, coincide with classical version

2. equals error exponent, just as its classical version

$$\sup\{b; P^n\{1|0\} \rightarrow 0\} = \underline{D}(\{\rho_0^n\}||\{\rho_1^n\})$$

$$\sup\{b; \limsup P^n\{1|0\} < 1\} = \bar{D}(\{\rho_0^n\}||\{\rho_1^n\})$$

3. If  $\rho_\theta^n = \rho_\theta^{\otimes n}$  both  $\bar{D}$  and  $\underline{D}$  coincide with  $D(\rho_0||\rho_1) := \text{tr}\rho_0 (\log\rho_0 - \log\rho_1)$

$$D(\rho_0||\rho_1) = \sup\{b; P^n\{1|0\} \rightarrow 0\} = \sup\{b; \limsup P^n\{1|0\} < 1\}$$

↑  
Had been known [日合Petz 80]

↑  
was not known [OgawaNagaoka 00]

# Another representation of $\overline{D}$ : smooth Renyi

In fact, spectrum-like quantity have been also used in computer science, for analysis of random number generation, security of cryptography (randomness extractor)

Its quantum version by R. Renner

$$D_{\max}(\rho_0 || \rho_1) = \min\{b; \rho_0 \leq e^b \rho_1\}$$

$$D_{\max}^\epsilon(\rho_0 || \rho_1) = \min\{D_{\max}(\rho_0' || \rho_1); \|\rho_0' - \rho_1\| \leq \epsilon\}$$

$$\lim_{\epsilon \downarrow 0} \limsup_{n \rightarrow \infty} \frac{1}{n} D_{\max}^\epsilon(\rho_0^n || \rho_1^n) = \overline{D}(\{\rho_0^n\} || \{\rho_1^n\})$$

[RennarDatta 05]

# Yet some more representations ...

$$\lim_{\epsilon \downarrow 0} \limsup_{n \rightarrow \infty} \frac{1}{n} D_{\max}^{\epsilon}(\rho_0^n || \rho_1^n) = \overline{D}(\{\rho_0^n\} || \{\rho_1^n\})$$

[RennarDatta 05]

$$D_{\max}^{\epsilon}(\rho_0 || \rho_1) = \min\{D_{\max}(\rho_0' || \rho_1); \|\rho_0' - \rho_1\| \leq \epsilon\}$$

$$\lim_{\epsilon \downarrow 0} \limsup_{n \rightarrow \infty} \frac{1}{n} D_{\alpha}^{\epsilon}(\rho_0^n || \rho_1^n) = \overline{D}(\{\rho_0^n\} || \{\rho_1^n\})$$

$$D_{\alpha}^{\epsilon}(\rho_0 || \rho_1) = \min\{D_{\alpha}(\rho_0' || \rho_1); \|\rho_0' - \rho_1\| \leq \epsilon\}$$

$D_{\alpha}(\rho_0 || \rho_1)$  = any CPTP-monotone quantum version of  $\alpha$ -Renyi relative entropy  
( $\alpha > 1$ )

$$\text{e.g. } \frac{1}{1-\alpha} \log \text{tr} \rho_0^{1-\alpha} \rho_1^{\alpha}$$

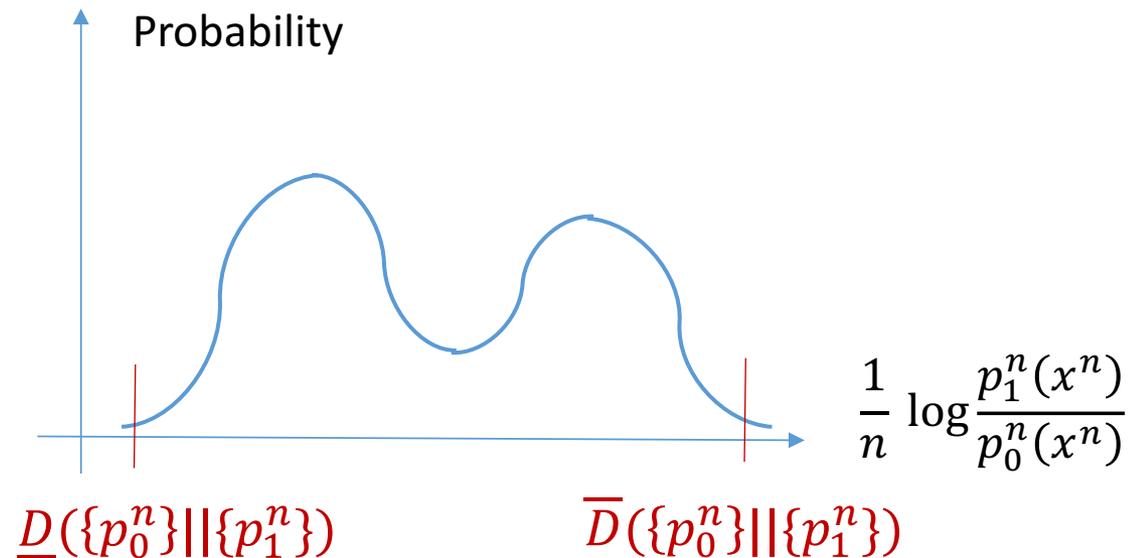
# on “n”

“n” is merely an index. Can be any number  
may not be system size....

It could be size of reservoir, for example

# When $\bar{D}$ and $D_\alpha (\rho_0' || \rho_1)$ coincide with D?

- Obviously, if iid.
- In classical case, when  $n$  is the system size, and the system is Markovian etc,  
(Maybe, when the systems are not globally correlated)



# An application : resolvability [Ogawa M]

$$x_1^n, x_2^n \cdots x_{L_n}^n \sim p^n$$

$\frac{1}{L_n} \sum_{i=1}^{L_n} \rho_{x_i^n}^n$  will be very close to  $E^n \rho_{x_1^n}^n$  if  $L_n$  is large enough.

How large is enough large ?

$$E^n \left\| \frac{1}{L_n} \sum_{i=1}^{L_n} \rho_{x_i^n}^n - E^n \rho_{x_1^n}^n \right\|_1 \rightarrow 0$$

holds if  $\limsup_{n \rightarrow \infty} \frac{1}{n} \log L_n \geq \sup_{x^n} \bar{D}(\{\rho_{x^n}^n\} || \{E^n \rho_{x_1^n}^n\})$

Putting altogether :

Divergence rate and Thermo-Majorization

# Classical case: $|\Theta| = 2, \mathbf{E}^n = \{p_0^n, p_1^n\}$

By Thermo-Majorization with error

$$\mathbf{E}^n = \{p_0^n, p_1^n\} \geq_{e_n} \mathbf{E}'^n = \{q_0^n, q_1^n\} \quad \lim e_{n,0} = 0, e_{n,1} = 0$$

$$(\Lambda(p_0^n) \approx q_0^n, \Lambda(p_1^n) = q_1^n)$$

$$\Leftrightarrow \forall b \sum_x (p_0^n(x) - e^{bn} p_1^n(x))_+ \geq \sum_x (q_0^n(x) - e^{bn} q_1^n(x))_+ + e_{n,0}$$

もしも寝ていなければ以下が見える筈

Thm

$$\begin{aligned} \mathbf{E}^n \geq_{e_n} \mathbf{E}'^n &\Leftrightarrow \underline{D}(\{p_0^n\} || \{p_1^n\}) \geq \bar{D}(\{q_0^n\} || \{q_1^n\}) \\ &\Rightarrow \bar{D}(\{p_0^n\} || \{p_1^n\}) \geq \bar{D}(\{q_0^n\} || \{q_1^n\}) \\ &\quad \underline{D}(\{p_0^n\} || \{p_1^n\}) \geq \underline{D}(\{q_0^n\} || \{q_1^n\}) \end{aligned}$$

$$\bar{D}(\{\rho_0^n\} || \{\rho_1^n\}) = \inf\{b; \lim_{n \rightarrow \infty} \text{tr}(\rho_0^n - e^{nb} \rho_1^n)_+ = 0\}$$

$$\underline{D}(\{\rho_0^n\} || \{\rho_1^n\}) = \sup\{b; \lim_{n \rightarrow \infty} \text{tr}(\rho_0^n - e^{nb} \rho_1^n)_+ = 1\}$$

# Classical case: $|\Theta| = 2$ , $E^n = \{p_0^n, p_1^n\}$

## Fact

$$E^n = \{p_0^n, p_1^n\} \geq_{e_n} E'^n = \{q_0^n, q_1^n\} \quad \liminf e_{n,0} < 2, e_{n,1} = 0$$

$$\Leftrightarrow \bar{D}(\{p_0^n\} || \{p_1^n\}) \geq \underline{D}(\{q_0^n\} || \{q_1^n\})$$

$\|\Lambda(p_0^n) - q_0^n\|_1 < 2$ , 完全に区別できるほどには悪くない

$$\Lambda(p_1^n) = q_1^n$$

$$\bar{D}(\{\rho_0^n\} || \{\rho_1^n\}) = \inf\{b; \lim_{n \rightarrow \infty} \text{tr}(\rho_0^n - e^{nb} \rho_1^n)_+ = 0\}$$

$$\underline{D}(\{\rho_0^n\} || \{\rho_1^n\}) = \sup\{b; \lim_{n \rightarrow \infty} \text{tr}(\rho_0^n - e^{nb} \rho_1^n)_+ = 1\}$$

# Standard form: $|\Theta| = 2$ , $E^n = \{\rho_0^n, \rho_1^n\}$

古典のときみたいに安直じゃないので頭捻る

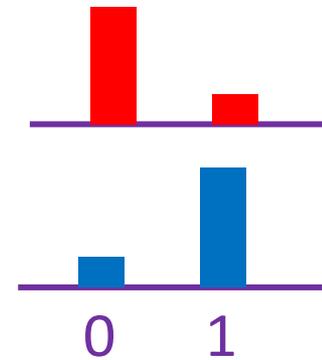
Standard form of binary experiments (dichotomies)

$\{r_0^n, r_1^n\}$  : probability distributions over  $\{0,1\}$

$$\lim r_0^n(0) = 1$$

$$\lim r_1^n(0) = 0, \text{ and exponentially}$$

$$r_1^n(0) \sim e^{-nb}$$



hypothesis test = conversion from  $E^n = \{\rho_0^n, \rho_1^n\}$  to  $\{r_0^n, r_1^n\}$

Among various conversions,  
Choose the one maximizing  $b$

$$\max b = \underline{D}(\{\rho_0^n\} || \{\rho_1^n\})$$

# Standard form: $|\Theta| = 2$ , $E^n = \{\rho_0^n, \rho_1^n\}$

It is natural to consider conversion from  $\{r_0^n, r_1^n\}$  to  $E^n = \{\rho_0^n, \rho_1^n\}$

**Lem** If  $c = \exp(-D_{\max}^\epsilon(\rho_0^n || \rho_1^n))$ , There is a states  $\tau_0, \tau_1$  with

$$\Leftrightarrow \|\tau_0 - \rho_0^n\|_1 \leq \epsilon \quad c\tau_0 + (1 - c)\tau_1 = \rho_1^n$$

**Proof : obvious from the def.**  $D_{\max}(\rho_0 || \rho_1) = \min\{b; \rho_0 \leq e^b \rho_1\}$

$$D_{\max}^\epsilon(\rho_0 || \rho_1) = \min\{D_{\max}(\rho_0' || \rho_1); \|\rho_0' - \rho_1\| \leq \epsilon\}$$

**strategy** By mixing  $\tau_0, \tau_1$  with probability  $r_\theta^n(0), r_\theta^n(1)$ , obtain  $\rho_0^n, \rho_1^n$

**Thm** There is a CPTP map  $\Lambda_n$  with  $\|\Lambda_n(r_0^n) - \rho_0^n\|_1 \rightarrow 0, \Lambda_n(r_1^n) = \rho_1^n$

if  $r_0^n(0) \rightarrow 0, r_1^n(0) = e^{-nb}$ ,

$$b > \lim_{\epsilon \downarrow 0} \limsup_{n \rightarrow \infty} \frac{1}{n} D_{\max}^\epsilon(\rho_0^n || \rho_1^n) = \bar{D}(\{\rho_0^n\} || \{\rho_1^n\})$$

$$|\Theta| = 2, \quad \mathbf{E}^n = \{\rho_0^n, \rho_1^n\}$$

$$\mathbf{E}^n \geq_{e_n} \mathbf{E}'^n \quad \lim e_{n,0} = 0, e_{n,1} = 0$$

$$(\Lambda(\rho_0^n) \approx \sigma_0^n, \quad \Lambda(\rho_1^n) = \sigma_1^n)$$

$$\mathbf{E}^n = \{\rho_0^n, \rho_1^n\} \rightarrow \{r_0^n, r_1^n\} \rightarrow \mathbf{E}'^n = \{\sigma_0^n, \sigma_1^n\}$$

Thm

$$\mathbf{E}^n \geq_{e_n} \mathbf{E}'^n \Leftrightarrow \underline{D}(\{\rho_0^n\} || \{\rho_1^n\}) \geq \bar{D}(\{\sigma_0^n\} || \{\sigma_1^n\})$$

$$\Rightarrow \bar{D}(\{\rho_0^n\} || \{\rho_1^n\}) \geq \bar{D}(\{\sigma_0^n\} || \{\sigma_1^n\})$$

$$\Rightarrow \underline{D}(\{\rho_0^n\} || \{\rho_1^n\}) \geq \underline{D}(\{\sigma_0^n\} || \{\sigma_1^n\})$$

If  $D = \bar{D} = \underline{D}$ , only have to compare  $D$

Can prove uniqueness of quantum extension of relative entropy, which is “smooth” and CPTP monotone

# Stability

Thm

If  $D^Q(\Lambda(\rho) || \Lambda(\sigma)) \leq D^Q(\rho || \sigma)$

$$D(\rho || \sigma) = \text{clim}_{\epsilon \downarrow 0} \lim_{n \rightarrow \infty} \frac{1}{n} \inf \{ D^Q(\rho' || \sigma^{\otimes n}); \|\rho' - \rho^{\otimes n}\|_1 \leq \epsilon \}$$

おしまい